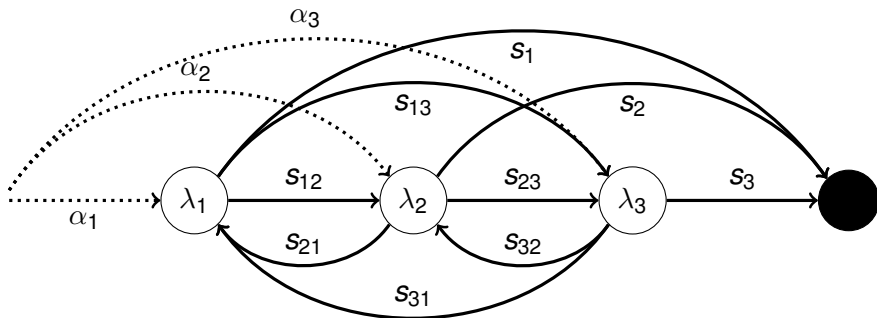




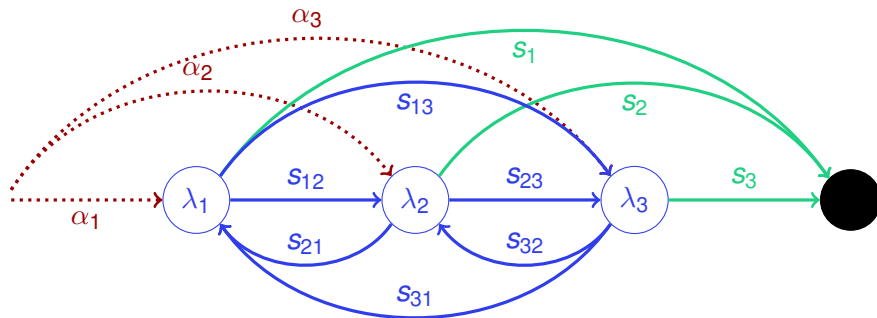
Clara Brimnes Gardner

# Phase-type Representations for Exponential Distributions

# Markov Jump Processes

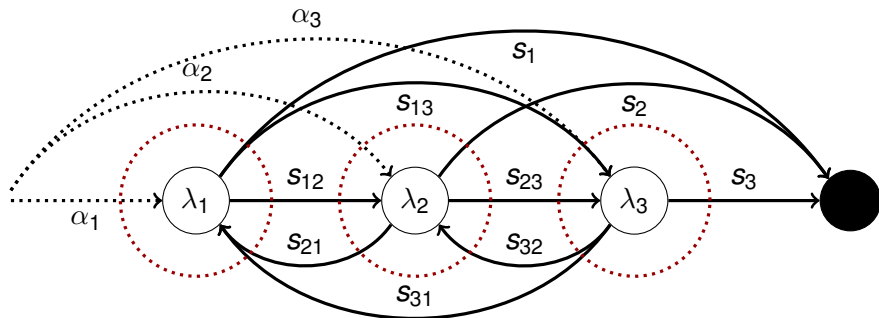


# Markov Jump Processes



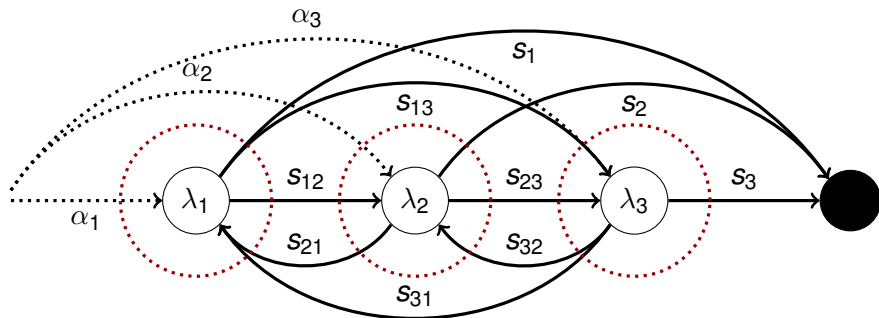
$$(\boldsymbol{\alpha}, \mathbf{S}) = \left( (\alpha_1 \quad \alpha_2 \quad \alpha_3), \begin{pmatrix} -\lambda_1 & s_{12} & s_{13} \\ s_{21} & -\lambda_2 & s_{23} \\ s_{31} & s_{32} & -\lambda_3 \end{pmatrix} \right) \text{ and } \mathbf{s} = \begin{pmatrix} s_1 \\ s_2 \\ s_3 \end{pmatrix} = -\mathbf{S}\mathbf{e}.$$

# Phase-type Distributions and the Exponential Distribution



$$Z_{1,i} \sim \exp(\lambda_1) \quad Z_{2,i} \sim \exp(\lambda_2) \quad Z_{3,i} \sim \exp(\lambda_3)$$

## Phase-type Distributions and the Exponential Distribution



$$Z_{1,i} \sim \exp(\lambda_1) \quad Z_{2,i} \sim \exp(\lambda_2) \quad Z_{3,i} \sim \exp(\lambda_3)$$

$$Y = \sum_{i=1}^p \sum_{j=1}^{N_i} Z_{i,j},$$

## Motivation

# Survival functions



## Motivation

# Survival functions



Exponential Distribution:

$$\mathbb{P}(Y > y) = \exp(-\lambda y),$$

Phase-type Distribution

$$\mathbb{P}(Y > y) = \alpha \exp(\mathbf{S}y) \mathbf{e}.$$

When do we have

$$\exp(-\lambda y) = \alpha \exp(\mathbf{S}y) \mathbf{e}?$$



The Jordan Normal Form

# The Jordan Normal Form



Figure: Picture from Wikipedia

## The Jordan Normal Form: General Expression

$$\mathbf{S} = \mathbf{T}\mathbf{J}\mathbf{T}^{-1} = (\mathbf{T}_1 \quad \mathbf{T}_2 \quad \cdots \quad \mathbf{T}_g) \begin{pmatrix} \mathbf{J}_1 & \mathbf{0} & \cdots & \mathbf{0} \\ \mathbf{0} & \mathbf{J}_2 & \cdots & \mathbf{0} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{0} & \mathbf{0} & \cdots & \mathbf{J}_g \end{pmatrix} \begin{pmatrix} \mathbf{U}_1 \\ \mathbf{U}_2 \\ \vdots \\ \mathbf{U}_g \end{pmatrix}$$

with

$$\mathbf{T}_i = (\mathbf{r}_{i,1} \quad \mathbf{r}_{i,2} \quad \cdots \quad \mathbf{r}_{i,b_i}), \quad \mathbf{J}_i = \begin{pmatrix} -\lambda_i & 1 & \cdots & 0 \\ 0 & -\lambda_i & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & -\lambda_i \end{pmatrix}, \quad \mathbf{U}_i = \begin{pmatrix} \mathbf{l}_{i,b_i} \\ \mathbf{l}_{i,b_i-1} \\ \vdots \\ \mathbf{l}_{i,1} \end{pmatrix}$$

### Example

$$\mathbf{S} = \begin{pmatrix} -4 & 1 & 2 & 0 \\ 0 & -2 & 0 & 1 \\ 0 & 0 & -3 & 3 \\ 0 & 0 & 0 & -3 \end{pmatrix},$$

The eigenvalues are  $-\lambda_1 = -2$ ,  $-\lambda_2 = -3$  and  $-\lambda_3 = -4$ .

### Example

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The eigenvalues are  $-\lambda_1 = -2$ ,  $-\lambda_2 = -3$  and  $-\lambda_3 = -4$ .

We can find three right eigenvectors: one for each eigenvalue

$$\mathbf{S}\mathbf{r}_1 = -\lambda_1\mathbf{r}_1 \Rightarrow \mathbf{r}_1 = (1/2 \ 1 \ 0 \ 0)',$$

$$\mathbf{S}\mathbf{r}_2 = -\lambda_2\mathbf{r}_2 \Rightarrow \mathbf{r}_2 = (2 \ 0 \ 1 \ 0)',$$

$$\mathbf{S}\mathbf{r}_3 = -\lambda_3\mathbf{r}_3 \Rightarrow \mathbf{r}_3 = (1 \ 0 \ 0 \ 0)'.$$

## The Jordan Normal Form: Example

We find a generalized eigenvector of order  $k = 2$

$$\mathbf{S} \cdot \mathbf{r}_{2,2} = -\lambda_2 \cdot \mathbf{r}_{2,2} + \mathbf{r}_2 \Rightarrow \mathbf{r}_{2,2} = \begin{pmatrix} -7/3 & -1/3 & 0 & 1/3 \end{pmatrix}$$

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Then we can write  $\mathbf{S}$  as

$$\mathbf{S} = \mathbf{T}\mathbf{J}\mathbf{T}^{-1} = (\mathbf{r}_{1,1} \quad \mathbf{r}_{2,1} \quad \mathbf{r}_{2,2} \quad \mathbf{r}_{3,1}) \begin{pmatrix} -\lambda_1 & 0 & 0 & 0 \\ 0 & -\lambda_2 & 1 & 0 \\ 0 & 0 & -\lambda_2 & 0 \\ 0 & 0 & 0 & -\lambda_3 \end{pmatrix} \begin{pmatrix} \mathbf{l}_{1,1} \\ \mathbf{l}_{2,2} \\ \mathbf{l}_{2,1} \\ \mathbf{l}_{3,1} \end{pmatrix}$$

Jordan block of order 2

## The Jordan Normal Form: Compact Notation

$$\mathbf{S} = \mathbf{T}\mathbf{J}\mathbf{T}^{-1} = \mathbf{T} \left( \sum_{i=1}^g ((-\lambda_i)\mathbf{E}_i + \mathbf{N}_i) \right) \mathbf{T}^{-1},$$

From previous example

$$\mathbf{E}_2 = \begin{pmatrix} 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 \end{pmatrix}, \quad \mathbf{N}_2 = \begin{pmatrix} 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 \end{pmatrix}$$

# The Jordan Normal Form and the Matrix Exponential

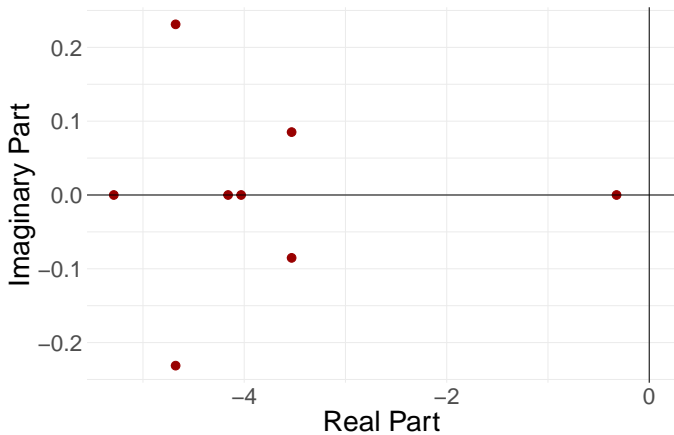
$$\begin{aligned}\exp(\mathbf{S}y) &= \mathbf{T} \exp(\mathbf{J}y) \mathbf{T}^{-1} \\ &= \sum_{j=1}^g \sum_{k=1}^{B_j} \mathbf{M}_{j,k} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_j y),\end{aligned}$$

with

$$\mathbf{M}_{j,k} = \mathbf{T} \left( \sum_{i \in A_j} \mathbf{E}_i \mathbf{N}_i^{k-1} \mathbf{E}_i \right) \mathbf{T}^{-1}.$$



# A look at the Eigenvalues



# The Survival Function

$$\mathbb{P}(Y > y) = \alpha \exp(\mathbf{S}y) \mathbf{e} = \alpha \sum_{i=1}^g \sum_{k=1}^{B_i} \mathbf{M}_{i,k} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_i y) \mathbf{e}$$

## The Survival Function

$$\begin{aligned}
 \mathbb{P}(Y > y) &= \alpha \exp(\mathbf{S}y) \mathbf{e} = \alpha \sum_{i=1}^g \sum_{k=1}^{B_i} \mathbf{M}_{i,k} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_i y) \mathbf{e} \\
 &= \mathbf{c}_{1,B_1} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_1 y) + \sum_{k=1}^{B_1-1} \mathbf{c}_{1,k} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_1 y) + \\
 &\quad \sum_{i=2}^g \sum_{k=1}^{B_i} \mathbf{c}_{i,k} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_i y)
 \end{aligned}$$

with

$$\mathbf{c}_{i,k} = \alpha \mathbf{M}_{i,k} \mathbf{e} = \alpha \mathbf{T} \left( \sum_{j \in A_i} \mathbf{E}_j \mathbf{N}_j^{k-1} \mathbf{E}_j \right) \mathbf{T}^{-1} \mathbf{e}.$$

# Paper from Bean and Green



PERGAMON

Mathematical and Computer Modelling 31 (2000) 31–46

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## When Is a MAP Poisson?

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**Abstract**—The departure process of a queue is important in the analysis of networks of queues, as it may be the arrival process to another queue in the network. A simple description of the departure process could enable a tractable analysis of a network, saving costly simulation or avoiding the errors of approximation techniques.

In a recent paper, Olivier and Walrand [1] conjectured that the departure process of a MAP/*PH*/1 queue is not a MAP unless the queue is a stationary *M*/*M*/1 queue. This conjecture was prompted by their claim that the departure process of an MMPP/*M*/1 queue is not a MAP unless the queue is a stationary *M*/*M*/1 queue. We note that their proof has an algebraic error, see [2], which leaves the above question of whether the departure process of an MMPP/*PH*/1 queue is a MAP, still open.

There is also a more fundamental problem with Olivier and Walrand's proof. In order to identify stationary *M*/*M*/1 queues, it is essential to be able determine from its generator when a stationary MAP is a Poisson process. This is not discussed in [1], nor does it appear to have been discussed elsewhere in the literature. This deficiency is remedied using ideas from nonlinear filtering theory, to give a characterisation as to when a stationary MAP is a Poisson process. © 2000 Elsevier Science Ltd. All rights reserved.

## Irreducible Generator Matrices

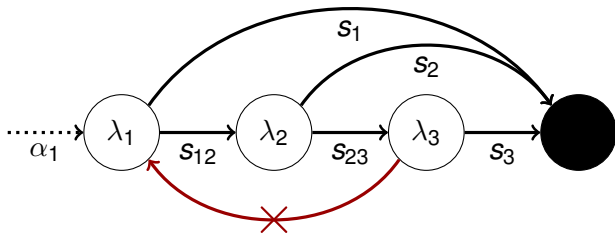
### Definition (Irreducible Generators)

A generator matrix,  $\mathbf{S}$  is said to be irreducible if **all states communicate**.

# Irreducible Generator Matrices

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A generator matrix,  $\mathbf{S}$  is said to be irreducible if **all states communicate**.



## Irreducible Representations

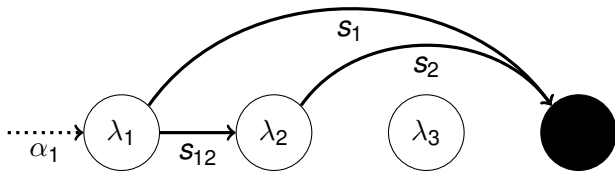
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A phase-type representation,  $(\alpha, \mathbf{S})$ , is said to be irreducible if there is a positive probability of visiting any of the transient states.

# Irreducible Representations

## Definition (Irreducible Representations)

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## The Transition Probability Matrix

The transition probabilities

$$\mathbb{P}(X_{y+t} = j \mid X_t = i) = p_{ij}(y), \quad t, y > 0,$$

are collected in the transition probability matrix

$$\{p_{ij}(y)\} = P(y) = \exp(\mathbf{S}y).$$

Using the Jordan normal form we obtain:

$$\exp(\mathbf{S}y) = \sum_{i=1}^g \sum_{k=1}^{B_i} \mathbf{M}_{i,k} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_i y)$$

## Probabilistic Argument: Non-negativity

$$\exp(\mathbf{S}y) = M_{1,B_1} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_1 y) + \sum_{k=1}^{B_1-1} M_{1,k} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_1 y) + \sum_{i=2}^g \sum_{k=1}^{B_i} M_{i,k} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_i y)$$

## Probabilistic Argument: Non-negativity

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Conclusion: Let  $\{m_{1,B_1}^{l,k}\}$  be the elements of  $M_{1,B_1}$ . Then

$$m_{1,B_1}^{l,k} \geq 0.$$

## Linear Algebra Argument: Positivity

The matrix  $M_{1,B_1}$  can be calculated as

$$\begin{aligned} M_{1,B_1} &= \mathbf{T} \left( \sum_{j \in A_i} \mathbf{E}_j \mathbf{N}_j^{k-1} \mathbf{E}_j \right) \mathbf{T}^{-1} \\ &= \mathbf{r}_{i,1} \mathbf{l}_{i,1}. \end{aligned}$$

But  $\mathbf{r}_{i,1} \neq \mathbf{0}$  and  $\mathbf{l}_{i,1} \neq \mathbf{0}$ .

## Linear Algebra Argument: Positivity

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But  $\mathbf{r}_{i,1} \neq \mathbf{0}$  and  $\mathbf{l}_{i,1} \neq \mathbf{0}$ .

Conclusion:  $M_{1,B_1}$  has at least one positive element.

## What About the Survival Function?

$$\begin{aligned}\mathbb{P}(Y > t + y) &= \alpha \exp(\mathbf{S}(t + y)) \mathbf{e} \\ &= \alpha \exp(\mathbf{S}t) \exp(\mathbf{S}(y)) \mathbf{e} \\ &= \mathbf{q}(t) \exp(\mathbf{S}y) \mathbf{e},\end{aligned}$$

where  $\mathbf{q}(t) = \alpha \exp(\mathbf{S}t)$ .

## What About the Survival Function?

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where  $\mathbf{q}(t) = \alpha \exp(\mathbf{S}t)$ .

If  $(\alpha, \mathbf{S})$  is an irreducible representation, then  $\mathbf{q}(t) > 0$ .

## Survival Function II

$$\begin{aligned}
 \mathbb{P}(y > t + y) &= \mathbf{q}(t) \exp(\mathbf{S}y) \mathbf{e} \\
 &= \mathbf{q}(t) M_{1, B_1} \mathbf{e} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_1 y) + \\
 &\quad \sum_{k=1}^{B_1-1} \mathbf{q}(t) M_{1, k} \mathbf{e} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_1 y) + \\
 &\quad \sum_{i=2}^g \sum_{k=1}^{B_i} \mathbf{q}(t) M_{i, k} \mathbf{e} \frac{y^{k-1}}{(k-1)!} \exp(-\lambda_i y)
 \end{aligned}$$



# Theorem

## Theorem (Exponential Phase-type Representations)

Let  $Y \sim PH(\alpha, \mathbf{S})$  such that the representation is irreducible.  $Y$  is exponentially distributed if and only if  $B_1 = 1$  and

$$c_{j,k} = \alpha \mathbf{T} \left( \sum_{i \in A_j} \mathbf{E}_i \mathbf{N}_i^{\nu-1} \mathbf{E}_i \right) \mathbf{T}^{-1} \mathbf{e} = 0.$$

for  $\{j, k\} \in \{2, \dots, g\} \times \{1, \dots, B_j\}$ .

## Theorem: Distinct Eigenvalues

### Theorem

Let  $Y \sim PH(\alpha, \mathbf{S})$  such that the representation is irreducible, and each Jordan block has a distinct eigenvalue.  $Y$  is exponentially distributed if and only if  $B_1 = 1$  and

$$\alpha \mathbf{r}_{j,\nu} = 0 \text{ or } \mathbf{l}_{j, B_j+1-\nu} \mathbf{e} = 0, \quad (1)$$

for all  $\{j, \nu\} \in \{2, 3, \dots, g\} \times \{1, 2, \dots, B_j\}$ .

## Example

Example: three distinct eigenvalues  $-\lambda_1$ ,  $-\lambda_2$  and  $-\lambda_3$  each with algebraic and geometric multiplicity :

$$\begin{aligned}\mathbb{P}(Y > y) &= \alpha \exp(\mathbf{S}y) \mathbf{e} \\ &= \alpha \mathbf{r}_1 \mathbf{l}_1 \mathbf{e} \exp(-\lambda_1 y) + \alpha \mathbf{r}_2 \mathbf{l}_2 \mathbf{e} \exp(-\lambda_2 y) + \alpha \mathbf{r}_3 \mathbf{l}_3 \mathbf{e} \exp(-\lambda_3 y)\end{aligned}$$

For  $Y$  to be exponentially distributed we require:

$$\alpha \mathbf{r}_2 = 0 \text{ or } \mathbf{l}_2 \mathbf{e} = 0,$$

$$\alpha \mathbf{r}_3 = 0 \text{ or } \mathbf{l}_3 \mathbf{e} = 0.$$

## PH-simplicity

### Definition (O' Cinneide)

Let  $\text{PH}(\mathbf{S})$  be the set of all phase-type distributions with a representation of the form  $(\alpha, \mathbf{S})$ . The generator  $\mathbf{S}$  is said to be PH-simple if each distribution in  $\text{PH}(\mathbf{S})$  has a unique representation of the form  $(\alpha, \mathbf{S})$ .

### Theorem (O' Cinneide)

*A PH-generator  $\mathbf{S}$  is PH-simple if and only if it has no left eigenvector orthogonal to  $\mathbf{e}$ .*

# Comparison

## Theorem

Let  $Y \sim PH(\alpha, \mathbf{S})$  such that the representation is irreducible, and each Jordan block has a distinct eigenvalue.  $Y$  is exponentially distributed if and only if  $B_1 = 1$  and

$$\alpha \mathbf{r}_{j,\nu} = 0 \text{ or } \mathbf{l}_{j,B_j+1-\nu} \mathbf{e} = 0, \quad (2)$$

for all  $\{j, \nu\} \in \{2, 3, \dots, g\} \times \{1, 2, \dots, B_j\}$ .

If  $\mathbf{S}$  is simple, then  $\mathbf{l}_{i,1} \mathbf{e} \neq 0$ . This implies that  $Y$  can only be exponentially distributed with  $\alpha \mathbf{r}_i = 0$ . This happens if  $\alpha$  is a left eigenvector.